

# Matrix Inverse Identities

Statistical Machine Learning

Veronica Burt

# Woodbury Identity

- We talked about the Woodbury Identity in class, which states for  $A$  and  $C$  nonsingular,

$$(A - BC^{-1}E)^{-1}BC^{-1} = A^{-1}B(C - EA^{-1}B)^{-1}$$

- Applying this identity to Ridge Regression, we saw

$$\hat{\beta}_{ridge} = (\mathbb{X}^T\mathbb{X} + \lambda I)^{-1}\mathbb{X}^T\mathbf{Y} = \mathbb{X}^T(\mathbb{X}\mathbb{X}^T + \lambda I)^{-1}\mathbf{Y}$$

- This results in the inversion of an  $n \times n$  matrix as opposed to a  $p \times p$  matrix, which can be much less expensive in terms of computation time in the big data setting.

# Discovering Matrix Inverse Formulas

- Once a matrix inverse formula is known, it is easy to check that it is true: we just multiply the two matrices together to verify that the result is the Identity Matrix.
- However, discovering the formulas is a much more difficult task.
- Many matrix inverse formulas were discovered by using partitioned (block) matrices.

- Note the following identity for  $A$  nonsingular:

$$\begin{bmatrix} A^{-1} & 0 \\ -VA^{-1} & I \end{bmatrix} \begin{bmatrix} A & U \\ V & D \end{bmatrix} = \begin{bmatrix} I & A^{-1}U \\ 0 & D - VA^{-1}U \end{bmatrix}$$

- Using the above equation, we see:

$$\begin{bmatrix} A & U \\ V & D \end{bmatrix}^{-1} = \begin{bmatrix} I & -A^{-1}U(D - VA^{-1}U)^{-1} \\ 0 & (D - VA^{-1}U)^{-1} \end{bmatrix} \begin{bmatrix} A^{-1} & 0 \\ -VA^{-1} & I \end{bmatrix} =$$

$$\begin{bmatrix} A^{-1} + A^{-1}U(D - VA^{-1}U)^{-1}VA^{-1} & -A^{-1}U(D - VA^{-1}U)^{-1} \\ -(D - VA^{-1}U)^{-1}VA^{-1} & (D - VA^{-1}U)^{-1} \end{bmatrix}$$

- This leads us to the conclusion that inverting a partitioned matrix leads to inverting the sum of two matrices.
- There are many versions of the previous equation, depending on which matrices we require to be nonsingular. For example, here is another version where  $A, B, U,$  and  $V$  are all nonsingular:

$$\begin{bmatrix} A & U \\ V & D \end{bmatrix}^{-1} = \begin{bmatrix} (A - UD^{-1}V)^{-1} & (V - DU^{-1}A)^{-1} \\ (U - AV^{-1}D)^{-1} & (D - VA^{-1}U)^{-1} \end{bmatrix}$$

# Linear Mixed Models

- Consider the linear mixed model,

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{u} + \boldsymbol{\epsilon},$$

where  $\mathbf{u}$  has dispersion matrix  $D$ , independent of  $\boldsymbol{\epsilon}$  which has dispersion matrix  $R$ .

- This results in  $\mathbf{Y}$  having the expected value of  $\mathbf{X}\boldsymbol{\beta}$  and covariance matrix  $(R + \mathbf{Z}D\mathbf{Z}^T)$ .
- To find the least squares estimate of  $\boldsymbol{\beta}$ , we must invert  $(R + \mathbf{Z}D\mathbf{Z}^T)$ , an  $n \times n$  matrix that often does not have a nice structure(it's normally large and nondiagonal).

- However, the Henderson equations give us

$$(R + ZDZ^T)^{-1} = R^{-1} - R^{-1}Z(Z^T R^{-1}Z + D^{-1})^{-1}Z^T R^{-1},$$

which requires us to invert  $R$ , and  $n \times n$  matrix that often has a nice structure,  $D$ , a  $q \times q$  matrix, and  $(Z^T R^{-1}Z + D^{-1})$ , which is a  $q \times q$  matrix.

- Depending on the size of  $n$  and  $q$  and the structure of  $D$  and  $R$ , we can choose the formula for  $\beta$  which minimizes computation time.

# Other Applications

- Intraclass correlation matrices
- Factor analysis
- Discriminant analysis
- Maximum likelihood estimation of variance components



## References

Henderson, H.V. and Searle, S. R.

*On Deriving the Inverse of a Sum of Matrices*

SIAM Review, Vol 23 No 1. 1981.